Circulating and Urinary Transforming Growth Factor β1, Amadori Albumin, and Complications of Type 1 Diabetes

The EURODIAB Prospective Complications Study

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OBJECTIVE — Transforming growth factor (TGF)-β1 is overexpressed in diabetes as a consequence of hyperglycemia and the creation of early glycated end products and may be responsible for the characteristic structural renal changes associated with diabetes. We sought to examine the role of both urinary and circulating TGF-β1 and its promotor Amadori albumin in the vascular complications of type 1 diabetes.

RESEARCH DESIGN AND METHODS — The present article reports on a nested case-control study from the EURODIAB Prospective Complications Study of Europeans with type 1 diabetes. Case subjects (n = 356) were all individuals with one or more complications of diabetes; control subjects (n = 189) were all individuals with no evidence of complications.

RESULTS — Urinary TGF-β1 and Amadori albumin were elevated in patients with micro- or macroalbuminuria. Standardized regression effects (SREs) for macroalbuminuria versus normoalbuminuria were 2.45 (95% CI 1.88–3.18, P = 0.0001 for urinary TGF-β1) and 1.67 (1.34–2.07, P = 0.001 for Amadori albumin). The SRE for urinary TGF-β1 remained statistically significant when adjusted for Hba1c, Amadori albumin, and blood pressure. Circulating TGF-β1 was elevated in individuals with proliferative retinopathy compared with individuals without retinopathy (SRE 1.29 [1.07–1.550], P = 0.007). This result was attenuated to 1.16 (0.95–1.43, P = 0.2) in the multivariate model, largely because of Hba1c.

CONCLUSIONS — Elevated levels of urinary TGF-β1 in macroalbuminuria were associated with elevations in Amadori albumin and Hba1c and also in blood pressure. In contrast, only circulating TGF-β1 was related to proliferative retinopathy, and Hba1c largely accounted for this. These findings may indicate novel pathways for understanding mechanisms and therapeutic interventions.

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improvements in glycemic control and the greater use of antihypertensive therapy should eventually have a beneficial impact on the incidence of severe nephropathy in type 1 diabetes. However, evidence that antihypertensive use in normotensive normoalbuminuric patients is of clinical value (1). Both microalbuminuria and macroalbuminuria significantly increase the risk of morbidity and mortality from coronary heart disease and are strong predictors of subsequent severe renal disease (2). Prevention of the early stages of diabetic renal disease and reduction in progression should now be priorities, but these steps require a more complete understanding of the etiology to identify suitable targets for intervention.

Diabetic nephropathy is characterized by hypertrophy of the glomerular and tubuloepithelial structures and thickening of the glomerular and tubular basement membrane, due largely to the effects of hyperglycemia (3). The cytokine transforming growth factor (TGF)-β1 appears to be a key mediator for these changes (4). TGF-β1 expression is enhanced in the presence of diabetes, either as a direct consequence of hyperglycemia or indirectly via the formation of early or advanced glycation end products (4). Hyperglycemia stimulates condensation reactions between glucose and proteins, and an early product of this reaction is an Amadori protein. Amadori albumin is the major early glycated protein, and such proteins may undergo further modifications to form advanced glycation end products.

TGF-β1 also upregulates the expression of a serine/threonine kinase, hSGK, which indirectly stimulates the renin angiotensin system (5). Angiotensin II is a potent stimulus for TGF-β1 (6) and may be the mechanism by which hypertension results in renal damage. Angiotensin II may also further explain why blockade of the renin angiotensin system is particularly effective in slowing progression of renal disease (7).

This defective extracellular matrix remodeling may also occur in other organs; therefore, it is reasonable to hypothesize that TGF-β1 may also be involved in the
other complications of diabetes. One previous study of type 2 diabetes indicated that this was indeed the case, but it is not clear whether these observations were confounded by albuminuria (8). We therefore examined the role of TGF-β1 and its possible inducers glucose and Amadori albumin in the complications of type 1 diabetes.

**RESEARCH DESIGN AND METHODS** — The EURODIAB Prospective Complications Study is a follow-up of the EURODIAB IDDM Complications Study (9). Baseline investigations (1988–1991) were performed on 3,250 men and women with type 1 diabetes drawn from 31 European centers. Sample selection was stratified by sex, age-group, and duration of diabetes to ensure sufficient representation in all categories. Type 1 diabetes was clinically defined as a diagnosis made before the age of 36 years, with a continuous need for insulin therapy within 1 year of diagnosis.

The follow-up (EURODIAB Prospective Complications Study) was performed on average 7–9 years later. Of the 3,250 patients, 1,880 (57.8%) returned for examination. Of these, full data on cardiovascular disease (CVD), microvascular complications, and neuropathy were available on 1,296 patients. Of those who were not reexamined, 437 had been recruited from centers not taking part in the follow-up, 101 had died, 465 provided morbidity data only, and 367 were untraced or otherwise unavailable.

All complications were again measured to a standard protocol (10). Two 24-h urine collections were performed (only one was performed at baseline) to measure the albumin excretion rate (AER). Aliquots were frozen and sent to London for analysis of urinary albumin with an immunoturbidimetric method using goat anti-human albumin antiserum (Sanofi Diagnostics Pasteur, Minneapolis, MN) and human serum albumin (ORHA 20/21 grade human serum albumin [HSA]; Behring Diagnostics, Hoechst, U.K.) standards (11). Retinopathy was assessed from retinal photographs (two fields per eye) according to the EURODIAB protocol (12). CVD was defined as a positive medical history, bypass graft surgery, and/or major Q waves on a centrally coded electrocardiogram.

Blood samples were taken, fasting if possible, for measurement of lipids and glycemic control. At a central laboratory, cholesterol and triglyceride levels were measured by enzymatic colorimetric tests (13), and HDL cholesterol was measured directly (14). HbA1c was measured by a latex enhanced turbidimetric immunoassay (Roche, Welwyn Garden City, U.K.). The reference range for this assay was 4.2–6.2%.

Total TGF-β1 was measured by an enzyme-linked immunosorbent assay (ELISA) development system (R&D Systems, Oxon, U.K.). To activate the latent plasma TGF-β1 to the immunoreactive form, 0.1 ml plasma was acid-activated by 0.1 ml 2.5 N acetic acid/10 mol/l urea for 10 min at room temperature and neutralized by 0.1 ml 2.7N NAOH/1 mol/l HEPES. For serum, 1 ml was activated by 0.2 ml 1 N HCl and neutralized by 0.2 ml 1.2 N NAOH/0.5 mol/l HEPES. Intra- and interassay variations were 5.0 and 7.0%, respectively. Urinary TGF-β1 was indexed to urinary creatinine excretion. Amadori albumin was determined in a competitive ELISA as recently described (15).

Briefly, each well was coated with 100 μl of 0.1 mg/ml glycated HSA as prepared by the incubation of HSA with 0.5 mol/l glucose for 4 weeks. This preparation was also used as standard in the competitive ELISA. Then, 50 μl of the antibody conjugated with biotin (1:2,000) and 50 μl of standard or a sample to be tested, diluted in PBS including 30 mg/ml HSA, was added to each well and incubated for 2 h. After three washes with PBS-Tween, the wells were incubated with streptavidin conjugated with biotin (1:2,000) and 50 μl tetramethylbenzidine. The reaction was stopped with 100 μl 0.5 mol/l H2SO4. The extinction at 450 nm was measured with a multichannel spectrophotometer (SLT Microplate Reader; Wilten Bioteknika, Etten-Leur, the Netherlands). Plasma levels were expressed as Amadori albumin units, and 1 unit was defined as the antibody-reactive material equivalent to 1 μg glycated HSA standard.

**Statistical analysis**

We used a nested case-control approach to maximize efficiency. Case subjects who had the greatest complication burden were selected to provide sufficient numbers for subgroup analyses. Control subjects who were completely free of complications were selected. Thus, case subjects were individuals with CVD or proliferative retinopathy or macroalbuminuria at follow-up and individuals with microalbuminuria and some degree of retinopathy (n = 356). Control subjects were individuals who had no evidence of CVD, retinopathy, or neuropathy and were normoalbuminuric at follow-up (n = 185). This selection allowed us to compare people with and without single or multiple complications. Case and control subjects were unmatched, so that the impact of key variables, such as age, could still be assessed, and any adjustments were taken care of at the analysis stage. Of these 541 individuals, sufficient sampling material was available for 535 for Amadori albumin, 534 for circulating TGF-β1, and 532 for urinary TGF-β1. These numbers allow us to detect a difference between groups of at worst one-third of an SD, with 90% power and 5% significance.

Key risk factors in case and control subjects were compared, reporting crude means and SDs or geometric means and 25th and 75th percentiles for skewed variables. Urinary and circulating TGF-β1 were log-transformed before analysis. Pearson correlation coefficients were first calculated for circulating and urinary TGF-β1 and Amadori albumin and other risk factors for complications. Then mean values of circulating and urinary TGF-β1 and Amadori albumin were estimated using ANOVA for each complication. Finally, standardized regression effects (SREs) were compared. These SREs were calculated by multiplying the β coefficient of each variable in a multivariate logistic regression model by its SD. This calculation allowed us to compare the impact of each factor directly with the other in standardized units. All SREs were first adjusted for age; adjustment for diabetes duration alone produced similar findings. There were no significant sex differences in these relationships; therefore, both sexes were combined, and no adjustment was made for sex.

**RESULTS**

**Differences between case and control subjects**

Case subjects were more likely to have an adverse risk factor profile and raised levels of urinary and circulating TGF-β1 and...
Table 1—Patient characteristics

<table>
<thead>
<tr>
<th>Characteristic</th>
<th>Case subjects</th>
<th>Control subjects</th>
</tr>
</thead>
<tbody>
<tr>
<td>n</td>
<td>356</td>
<td>185</td>
</tr>
<tr>
<td>Age (years)</td>
<td>41.7 ± 10.6</td>
<td>36.1 ± 8.1</td>
</tr>
<tr>
<td>Diabetes duration (years)</td>
<td>25.0 ± 9.0</td>
<td>15.5 ± 7.0</td>
</tr>
<tr>
<td>HbA1c (%)</td>
<td>9.0 ± 1.6</td>
<td>7.7 ± 1.3</td>
</tr>
<tr>
<td>Systolic blood pressure (mmHg)</td>
<td>127 ± 21.5</td>
<td>115 ± 13.7</td>
</tr>
<tr>
<td>Diastolic blood pressure (mmHg)</td>
<td>75 ± 11.7</td>
<td>74 ± 10.9</td>
</tr>
<tr>
<td>AER (µg/min)</td>
<td>65.4 (9.21, 372)</td>
<td>6.2 (4.5, 8.4)</td>
</tr>
<tr>
<td>Amadori albumin (units/ml)</td>
<td>48.1 ± 13.7</td>
<td>42.7 ± 13.1</td>
</tr>
<tr>
<td>Urinary TGF-β1/creatinine (pg/mmol)</td>
<td>3.71 (2.34, 5.26)</td>
<td>2.36 (1.55, 3.46)</td>
</tr>
<tr>
<td>Circulating TGF-β1 (ng/ml)</td>
<td>6.17 (3.71, 9.00)</td>
<td>5.52 (3.29, 8.08)</td>
</tr>
<tr>
<td>Cholesterol (mmol/l)</td>
<td>5.5 ± 1.24</td>
<td>5.0 ± 1.10</td>
</tr>
<tr>
<td>HDL cholesterol (mmol/l)</td>
<td>1.60 ± 0.43</td>
<td>1.67 ± 0.45</td>
</tr>
<tr>
<td>Triglyceride (mmol/l)</td>
<td>1.23 (0.85, 1.58)</td>
<td>0.84 (0.66, 1.08)</td>
</tr>
<tr>
<td>Waist (cm)</td>
<td>86.0 ± 11.9</td>
<td>83.0 ± 10.9</td>
</tr>
<tr>
<td>Height (cm)</td>
<td>168 ± 9.4</td>
<td>170 ± 9.6</td>
</tr>
<tr>
<td>Weight (kg)</td>
<td>70.1 ± 12.5</td>
<td>69.3 ± 11.9</td>
</tr>
<tr>
<td>Waist-to-hip ratio</td>
<td>0.89 ± 0.13</td>
<td>0.88 ± 0.16</td>
</tr>
</tbody>
</table>

Data are means ± SD or geometric means (25th and 75th percentiles) for log-transformed data.

Amadori albumin than control subjects (Table 1).

Correlations with other risk markers of complications

Urinary and circulating TGF-β1 was not correlated (coefficient 0.05, P = 0.3, Table 2). Urinary TGF-β1 was correlated with diabetes duration, HbA1c, systolic pressure, AER, Amadori albumin, and lipids and was correlated negatively with measures of body size. Circulating TGF-β1 was only correlated with HbA1c and Amadori albumin. Amadori albumin, as anticipated, was correlated with HbA1c, AER, and lipids and was correlated negatively with estimates of body size.

TGF-β1, Amadori albumin, and diabetic nephropathy

Mean urinary TGF-β1 and Amadori albumin were elevated in microalbuminuric and macroalbuminuric patients (Table 3). However, circulating TGF-β1 was only significantly raised in macroalbuminuric patients compared with normoalbuminuric patients.

The strength of the association between TGF-β1 and macroalbuminuria was second only to systolic blood pressure in the univariate analyses (Table 4, SRE 2.45 vs. 3.18). All univariate associations were attenuated in bivariate models. For example, the SRE for TGF-β1 was attenuated from 2.45 to 2.02 with the addition of HbA1c, and the univariate association between HbA1c and macroalbuminuria was attenuated from 2.27 to 1.91 in the same model. In the multivariate model, systolic blood pressure had an almost unaltered impact on macroalbuminuria compared with the univariate data, indicating that the impact of blood pressure on albuminuria is likely to be independent of the consequences of hyperglycemia.

Table 2—Correlation coefficients between urinary and circulating TGF-β1 and Amadori albumin and risk factors for microvascular complications

<table>
<thead>
<tr>
<th></th>
<th>TGF-β1 urine</th>
<th>TGF-β1 circulating</th>
<th>Amadori albumin</th>
</tr>
</thead>
<tbody>
<tr>
<td>n</td>
<td>532</td>
<td>534</td>
<td>535</td>
</tr>
<tr>
<td>Age</td>
<td>0.07</td>
<td>−0.03</td>
<td>0.09 (0.04)</td>
</tr>
<tr>
<td>Diabetes duration</td>
<td>0.14 (0.002)</td>
<td>−0.01</td>
<td>0.12 (0.005)</td>
</tr>
<tr>
<td>HbA1c</td>
<td>0.32 (0.0001)</td>
<td>0.15 (0.0004)</td>
<td>0.43 (0.0001)</td>
</tr>
<tr>
<td>Systolic blood pressure</td>
<td>0.13 (0.003)</td>
<td>0.05</td>
<td>0.09 (0.03)</td>
</tr>
<tr>
<td>Diastolic blood pressure</td>
<td>0.02</td>
<td>0.08</td>
<td>−0.02</td>
</tr>
<tr>
<td>AER</td>
<td>0.37 (0.0001)</td>
<td>0.07</td>
<td>0.22 (0.0001)</td>
</tr>
<tr>
<td>Amadori albumin</td>
<td>0.26 (0.0001)</td>
<td>0.21 (0.0001)</td>
<td>—</td>
</tr>
<tr>
<td>TGF-β1 urine</td>
<td>0.06</td>
<td>—</td>
<td>0.26 (0.0001)</td>
</tr>
<tr>
<td>TGF-β1 circulating</td>
<td>0.16 (0.0003)</td>
<td>0.08</td>
<td>0.16 (0.0002)</td>
</tr>
<tr>
<td>Cholesterol</td>
<td>0.07</td>
<td>0.04</td>
<td>0.16 (0.0002)</td>
</tr>
<tr>
<td>HDL cholesterol</td>
<td>0.14 (0.01)</td>
<td>−0.03</td>
<td>0.13 (0.02)</td>
</tr>
<tr>
<td>Triglyceride†</td>
<td>−0.10 (0.02)</td>
<td>0.00</td>
<td>−0.07</td>
</tr>
<tr>
<td>Waist</td>
<td>−0.20 (0.0001)</td>
<td>−0.08</td>
<td>−0.13 (0.003)</td>
</tr>
<tr>
<td>Height</td>
<td>−0.18 (0.0001)</td>
<td>−0.08 (0.05)</td>
<td>−0.11 (0.01)</td>
</tr>
<tr>
<td>Weight</td>
<td>−0.09 (0.03)</td>
<td>0.05</td>
<td>−0.06</td>
</tr>
</tbody>
</table>

Data in parentheses are P values. *Urinary TGF-β1 indexed to creatinine; †Fasting triglyceride only restricted to n = 332 for urinary TGF-β1, n = 334 for circulating TGF-β1, and n = 335 for Amadori albumin.
retinopathy, differed little from their univariate associations, but the relationship for circulating TGF-β1 and Amadori albumin became statistically nonsignificant.

### TGF-β1, Amadori albumin, and CVD

There was no association between Amadori albumin and either urinary or circulating TGF-β1 and CVD (data not shown).

#### TGF-β1 levels and ACE inhibitor use

Circulating TGF-β1 was higher in individuals on ACE inhibitors than in individuals not on any antihypertensive therapy (3.67 vs. 2.92 ng/ml, \(P = 0.001\)). A nonsignificant difference was observed for urinary TGF-β1 (5.99 vs. 5.88 pg/mmol, \(P = 0.9\)). These associations may be confounded by nephropathy status. Strikingly, elevated levels of circulating TGF-β1 were observed in individuals on ACE inhibitors in the normoalbuminuric and microalbuminuric range (5.87 vs. 5.58 and 7.46 vs. 6.05 ng/ml, respectively), whereas levels were lower in individuals on ACE inhibitors who had macroalbuminuria (5.58 vs. 7.03 ng/ml, respectively; \(P = 0.04\) for interaction). Similar differences were observed for urinary TGF-β1. Mean systolic blood pressure was 13 mmHg higher in normoalbuminuric patients on ACE inhibitors than in individuals not on treatment; this difference was 8 mmHg in microalbuminuric patients and 5 mmHg in macroalbuminuric patients. Adjustment for these blood pressure differences largely attenuated differences in both circulating and urinary TGF-β1 in normoalbuminuric and microalbuminuric patients but not in macroalbuminuric patients.

### CONCLUSIONS

We show that both urinary TGF-β1 and Amadori albumin are elevated in microalbuminuric and macroalbuminuric patients with type 1 diabetes. Separately, Amadori albumin and HbA1c accounted for a similar proportion of the association between TGF-β1 and albuminuria, but their joint impact attenuated this association still further. This result indicates that HbA1c effects TGF-β1 via mechanisms in addition to its impact on Amadori albumin and on albuminuria via means that are not exclusive to TGF-β1. This is not unexpected because it is likely that several other mechanisms account for the impact of hyperglycemia on diabetic renal disease.

### Table 3—Mean age-adjusted values of circulating and urinary TGF-β1 and Amadori albumin by albuminuria and retinopathy

<table>
<thead>
<tr>
<th></th>
<th>n</th>
<th>Urinary TGF-β1/creatinine (pg/mmol)</th>
<th>Circulating TGF-β1 (ng/ml)</th>
<th>Amadori albumin (units/ml)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Normoalbuminuria</td>
<td>326</td>
<td>2.6 (2.4–2.8)</td>
<td>5.6 (5.2–6.1)</td>
<td>43.2 (41.8–44.6)</td>
</tr>
<tr>
<td>Microalbuminuria</td>
<td>81</td>
<td>3.7 (3.2–4.3)</td>
<td>6.4 (5.5–7.5)</td>
<td>47.8 (44.9–50.6)</td>
</tr>
<tr>
<td>Macroalbuminuria</td>
<td>126</td>
<td>4.6 (4.1–5.2)</td>
<td>6.2 (5.5–7.0)</td>
<td>50.3 (48.1–52.6)</td>
</tr>
<tr>
<td>P for macro- vs. normoalbuminuria</td>
<td>—</td>
<td>0.001</td>
<td>0.001</td>
<td></td>
</tr>
<tr>
<td>P for micro- vs. normoalbuminuria</td>
<td>—</td>
<td>0.001</td>
<td>0.5</td>
<td></td>
</tr>
<tr>
<td>No retinopathy</td>
<td>235</td>
<td>2.5 (2.3–2.7)</td>
<td>5.7 (5.2–6.2)</td>
<td>43.5 (41.8–45.2)</td>
</tr>
<tr>
<td>Background retinopathy</td>
<td>146</td>
<td>3.6 (3.2–4.1)</td>
<td>5.4 (4.8–6.1)</td>
<td>45.2 (43.0–47.3)</td>
</tr>
<tr>
<td>Proliferative retinopathy</td>
<td>154</td>
<td>3.8 (3.4–4.3)</td>
<td>6.7 (6.0–7.4)</td>
<td>49.3 (47.2–51.4)</td>
</tr>
<tr>
<td>P for trend</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
</tr>
<tr>
<td>P for proliferative vs. none</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
</tr>
<tr>
<td>P for proliferative vs. none, adjusted for AER</td>
<td>—</td>
<td>0.06</td>
<td>0.02</td>
<td>0.0001</td>
</tr>
</tbody>
</table>

Data are means (95% CI) unless otherwise indicated.

### Table 4—(SREs) for macroalbuminuria versus normoalbuminuria by key risk factors

<table>
<thead>
<tr>
<th></th>
<th>SRE</th>
<th>95% CI</th>
<th>(P)</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Univariate</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>HbA1c</td>
<td>2.27</td>
<td>1.79–2.88</td>
<td>0.0001</td>
</tr>
<tr>
<td>Systolic blood pressure</td>
<td>3.18</td>
<td>2.37–4.27</td>
<td>0.0001</td>
</tr>
<tr>
<td>TGF-β1 urine</td>
<td>2.45</td>
<td>1.88–3.18</td>
<td>0.0001</td>
</tr>
<tr>
<td>Amadori albumin</td>
<td>1.67</td>
<td>1.34–2.07</td>
<td>0.0001</td>
</tr>
<tr>
<td><strong>Bivariate</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Systolic blood pressure and TGF-β1 urine</td>
<td>3.18</td>
<td>2.32–4.37</td>
<td>0.0001</td>
</tr>
<tr>
<td>TGF-β1 urine</td>
<td>2.31</td>
<td>1.76–3.04</td>
<td>0.0001</td>
</tr>
<tr>
<td>HbA1c and TGF-β1 urine</td>
<td>1.91</td>
<td>1.49–2.45</td>
<td>0.0001</td>
</tr>
<tr>
<td>HbA1c</td>
<td>2.02</td>
<td>1.54–2.65</td>
<td>0.0001</td>
</tr>
<tr>
<td>Amadori albumin and TGF-β1 urine</td>
<td>1.42</td>
<td>1.13–1.80</td>
<td>0.003</td>
</tr>
<tr>
<td>TGF-β1 urine</td>
<td>2.26</td>
<td>1.73–2.94</td>
<td>0.0001</td>
</tr>
<tr>
<td>HbA1c and Amadori albumin</td>
<td>2.10</td>
<td>1.59–2.64</td>
<td>0.0001</td>
</tr>
<tr>
<td>Amadori albumin</td>
<td>1.21</td>
<td>0.94–1.54</td>
<td>0.1</td>
</tr>
<tr>
<td><strong>Multivariate, all variables in model</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Systolic blood pressure</td>
<td>3.24</td>
<td>2.32–4.52</td>
<td>0.0001</td>
</tr>
<tr>
<td>TGF-β1 urine</td>
<td>1.86</td>
<td>1.40–2.48</td>
<td>0.0001</td>
</tr>
<tr>
<td>Amadori albumin</td>
<td>1.15</td>
<td>0.87–1.54</td>
<td>0.3</td>
</tr>
<tr>
<td>HbA1c</td>
<td>1.87</td>
<td>1.38–2.53</td>
<td>0.0001</td>
</tr>
</tbody>
</table>

All SREs have been adjusted for age.
Cell culture studies show that a high glucose medium can enhance autocrine production and activation of TGF-$\beta 1$, and TGF-$\beta 1$ in turn can enhance glucose uptake through fibroblasts (16), indicating an amplification of the effect of hyperglycemia. In human studies of type 2 diabetes, both urinary and circulating TGF-$\beta 1$ have been found to be raised in individuals with proteinuria but not in individuals with microalbuminuria compared with control subjects (17). However, these previous studies were relatively small and may not have been powered to detect a difference for earlier stages of disease. Further, these studies have not generally accounted for the confounding effects of other risk factors and complications. Few data exist for type 1 diabetes and have so far not indicated elevations of urinary TGF-$\beta 1$ with albuminuria as we do here (18).

We also show that the association between Amadori albumin and macroalbuminuria could largely be explained by glycemic control because the univariate SRE was attenuated from 1.67 to 1.21 in the bivariate model with HbA$_{1c}$, with little attenuation to 1.15 in the multivariate model, indicating that much of the mechanisms of HbA$_{1c}$ and Amadori albumin on diabetic nephropathy are shared and in part working through TGF-$\beta 1$. There is now good mechanism evidence demonstrating the role of Amadori albumin in the etiology of diabetic nephropathy. Mesangial cells can recognize and bind Amadori albumin in a dose-response fashion, and animal studies show that infusion of Amadori albumin produces renal glomerular changes that are similar to those observed in diabetes (19,20). That this nephropathic effect is mediated via TGF-$\beta 1$ is shown by cell culture studies, in which incubation with glycated albumin resulted in overexpression of TGF-$\beta 1$ mRNA levels (21). Administration of anti-glycated albumin antibody in animal models reduces plasma levels of glycated albumin, reduces AER, and reverses glomerular pathology (22). Human studies have been conflicting, perhaps in part because of small numbers and confounding by other risk factors, but one of the largest previous studies clearly indicated that Amadori albumin levels were elevated in individuals with nephropathy compared with individuals without nephropathy (15,23–25).

Interestingly, the inclusion of both Amadori albumin and HbA$_{1c}$ in our models could only account for around one-quarter of the association between TGF-$\beta 1$ and albuminuria, indicating that other mechanisms shape the association between TGF-$\beta 1$ and albuminuria. A clear candidate is hypertension, which alone was the strongest correlate of macroalbuminuria. Certainly the inclusion of blood pressure with TGF-$\beta 1$ had the same attenuating effect on the risk estimate of TGF-$\beta 1$ as Amadori albumin did, suggesting that raised blood pressure, an important risk factor for albuminuria, was in part working through TGF-$\beta 1$ to mediate its effects; however, clearly from the multivariate analyses, this was not the sole, or even the most important, mechanism. This finding can be accounted for by the observation of a stimulatory effect of TGF-$\beta 1$ on the renin angiotensin system (6) and the beneficial effect of angiotensin II receptor blockers on TGF-$\beta 1$ levels and renal pathology (26). Other mechanisms not studied here include activation of protein kinase C, other growth factors such as vascular endothelial growth factor and insulin-like growth factor, oxidative stress, and intraglomerular hemodynamic factors (27–29).

Interestingly, urinary TGF-$\beta 1$ was only lower in macroalbuminuric patients on ACE inhibitors compared with individuals on no treatment. In the normoalbuminuric and microalbuminuric categories, there appeared to be little difference or even a tendency to higher levels in individuals on ACE inhibitors. This was in part accounted for by the greater blood pressure differentials at lower albuminuria levels between individuals on treatment and those not on treatment. A separate explanation is that ACE inhibi-

### Table 5—Univariate and multivariate associations between risk factors and proliferative retinopathy (SREs)

<table>
<thead>
<tr>
<th>Risk Factor</th>
<th>SRE</th>
<th>95% CI</th>
<th>P</th>
</tr>
</thead>
<tbody>
<tr>
<td>Amadori albumin</td>
<td>1.45</td>
<td>1.20–1.75</td>
<td>0.0001</td>
</tr>
<tr>
<td>Circulating TGF-$\beta 1$</td>
<td>1.29</td>
<td>1.07–1.55</td>
<td>0.007</td>
</tr>
<tr>
<td>Systolic blood pressure</td>
<td>1.69</td>
<td>1.37–2.10</td>
<td>0.0001</td>
</tr>
<tr>
<td>HbA$_{1c}$</td>
<td>2.07</td>
<td>1.68–2.56</td>
<td>0.0001</td>
</tr>
</tbody>
</table>

All SREs have been adjusted for age.

For bivariate analysis, we used a threshold of $P<0.05$. Univariate analysis was performed with SREs for all risk factors, and multivariate analysis with the SRE for all risk factors except those in the same category.
tion suppresses TGF-β1 receptor expression, which could reduce TGF-β1 signaling and thus TGF-β1-mediated effects (30).

In contrast, circulating TGF-β1 was only elevated in the presence of macroalbuminuria, and not microalbuminuria, compared with normoalbuminuria, indicating that it may be less sensitive to renal changes. However, clear differences were observed between proliferative retinopathy and no retinopathy in circulating TGF-β1, which unlike differences observed for urinary TGF-β1, could not be accounted for by confounding with AER. The impact of TGF-β1 on proliferative retinopathy remained largely unexplained by elevated blood pressure but could be largely accounted for by HbA1c. Others (8) have indicated that circulating TGF-β1 is elevated in nonproliferative retinopathy in type 2 diabetes. By inhibiting replacement of damaged cells, TGF-β1 may contribute to capillary and pericyte loss in diabetic retinopathy (31).

In contrast though, whereas total vitreous TGF-β1 did not differ between individuals with proliferative retinopathy compared with control subjects, latent TGF-β1 in plasma samples appears to be relatively small (36), and because the protocol for sample collection and handling was identical across sites, any bias is likely to be nondifferential.

In conclusion, we demonstrate that urinary TGF-β1 is closely correlated with the degree of albuminuria in type 1 diabetes and confirm that this effect is largely accounted for by hyperglycemia, Amadori albumin, and blood pressure. Circulating TGF-β1, on the other hand, is more closely correlated with retinopathy, and its effect does not appear to be so closely associated with Amadori albumin. These findings underline new possibilities for both the monitoring and treatment of microvascular complications of type 1 diabetes and deserve further exploration.

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APPENDIX

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